### Class 6: Bayesian Methods

In this class we will review Bayesian likelihood methods for solving statistical problems, determining the posterior probabilities of model parameters, and selecting between two models

## Class 6: Bayesian Methods

At the end of this class you should be able to ...

- ... understand the application of Bayes' theorem in modelfitting and the role of priors
- ... obtain parameter values and confidence ranges via likelihood methods
- ... search parameter space with MCMC algorithms
- ... apply model selection tests using the Bayes factor or Akaike information criteria

• Recall from Class 1 that Bayesian statistics is a framework that allows us to assign probabilities to a model



- It makes use of conditional probabilities, P(A|B), meaning *"the probability of A on the condition that B has occurred"*
- Remember that  $P(A|B) \neq P(B|A)$  in general!

 An important role in Bayesian statistics is played by Bayes' theorem, which can be derived from elementary probability:

$$P(A|B) = \frac{P(B|A) P(A)}{P(B)}$$

• Small print: this formula can be derived by just writing down the joint probability of both A and B in 2 ways:

$$P(A \cap B) = P(A|B) P(B) = P(B|A) P(A)$$

- The chance of a certain medical test being positive is 90%, if a patient has disease *D*. 1% of the population have the disease, and the test records a false positive 5% of the time. If you receive a positive test, what is your probability of having *D*?
- We are told: P(+|D) = 0.9, P(D) = 0.01, P(+|no D) = 0.05
- We want to know: P(D|+)

• Bayes' Theorem: 
$$P(D|+) = \frac{P(+|D) P(D)}{P(+)} = \frac{P(+|D) P(D)}{P(+|D) P(D) + P(+|no D) P(no D)}$$

- Substituting in the data:  $P(D|+) = \frac{0.9 \times 0.01}{0.9 \times 0.01 + 0.05 \times 0.99} = 0.15$
- Interpretation: although the test is correct 90% of the time, the probability of having D after a positive test is only 15%. This is because only a small fraction of the population have the disease.

- A **Frequentist** might argue *"either the person has the disease* or not it is meaningless to apply probability in this way"
- A **Bayesian** might argue "there is a prior probability of 1% that the person has the disease. This probability should be updated in the light of the new data using Bayes' theorem"



• Bayes' theorem can be usefully re-written for science as:



# Role of the prior

- Bayesian statistics cannot determine probabilities of a model without assigning a prior probability
- The importance of the prior probability is both the strong and weak point of Bayesian statistics
- A Bayesian might argue: "the prior probability is a logical necessity when assessing the probability of a model. It should be stated, and if it's unknown you can use an uninformative (wide) prior"
- A Frequentist might argue "setting the prior is subjective two experiments could use the same data to come to two different conclusions, just by taking different priors"

# Role of the prior

• Let's take the example of fitting a parameter *a* to some data. Bayes' Theorem now reads:

 $P(a|\text{data}) \propto P(\text{data}|a) P(a)$ 

- We do not need the denominator, since we will normalize the posterior P(a|data) such that  $\int P(a|\text{data}) da = 1$
- In the absence of other information, a **uniform (or constant) prior** is often assumed for P(a). This is effectively equivalent to the **fitting range** of a parameter
- Assuming Gaussian variables, the **likelihood** P(data|a) is:



Hence:  $P(a|\text{data}) \propto e^{-\chi^2/2}$ 

### Posteriors and confidence limits

 We can use the posterior probability distribution P(a) to determine summary statistics and confidence intervals for the parameter a:

• Mean: 
$$\mu_a = \int_{-\infty}^{\infty} a P(a) da$$

• Variance:

$$\sigma_a{}^2 = \int_{-\infty}^{\infty} (a - \mu_a)^2 P(a) \, da$$

 [Small print: only if the probability distribution is Gaussian is the mean equal to the best-fitting value, and the standard deviation equal to the 68% confidence region]



### Posteriors and confidence limits

• For a general probability distribution, we can determine the confidence intervals by integration:



# Marginalization

- Now suppose we have determined the 2D posterior probability distribution of a 2-parameter fit,  $P_{2D}(a, b)$
- What is the probability distribution for parameter a, considering all possible values of parameter b? This is known as marginalization of parameter b
- Marginalization can be performed by summing (integrating) over one axis of the probability distribution:

$$P_{1D}(a) = \sum_{b} P_{2D}(a,b)$$

• [Small print: if  $P_{2D}(a, b)$  is normalized, then  $P_{1D}(a)$  will also be normalized]

• Let's apply these methods to our example from Class 3, fitting a straight line y = ax + b to some data...



• We determine the values of  $\chi^2$  over a grid of (a, b) and convert to 2D probability  $P(a, b) \propto e^{-\chi^2/2}$ 



• Then we marginalize to obtain the posterior probability distributions for each parameter, P(a) and P(b) ...



• By integrating under these distributions, we identify the 68% confidence regions ...



# Supernova cosmology (continued)

- Let's return to the same **supernova distance-redshift dataset** we were using in Class 3:
- Convert the  $\chi^2$  values into a joint 2D probability distribution in  $(\Omega_m, \Omega_\Lambda)$
- Marginalize this probability distribution to obtain the 1D posterior probability distributions for  $\Omega_m$  and  $\Omega_\Lambda$
- Determine the 68% confidence regions for  $\Omega_m$  and  $\Omega_\Lambda$



# Monte Carlo Markov Chains

- The grid method becomes inefficient as the number of parameters increases. A powerful alternative is to generate a Monte Carlo Markov Chain (MCMC) in the parameter space
- There are various algorithms to do this such as python *emcee* (we won't go into details here), but the end result is a "chain" (distribution of parameter values) which **samples the underlying probability distribution**



## Monte Carlo Markov Chains

• Here is a worked example of using python's emcee algorithm to sample the probability distribution of the straight-line fit:



# Supernova cosmology (continued)

- Let's return to the same **supernova distance-redshift dataset** again:
- Run an MCMC analysis for parameters  $(\Omega_m, \Omega_\Lambda)$
- Determine the 68% confidence regions for  $\Omega_m$  and  $\Omega_\Lambda$



### Model selection

- Since Bayesian statistics is related to the probability of models, it allows us to perform **model selection**
- A common example: how many model parameters does a dataset justify including in a fit?



### Model selection

 In general, given models M<sub>1</sub> (parameter p<sub>1</sub>) and M<sub>2</sub> (parameter p<sub>2</sub>) and a dataset D, we can determine the Bayes factor:

$$K = \frac{P(M_1|D)}{P(M_2|D)} = \frac{\int dp_1 P(D|p_1) P(p_1)}{\int dp_2 P(D|p_2) P(p_2)}$$

• The size of *K* quantifies how strongly we can prefer one model to another, e.g. the **Jeffreys scale**:

K	Strength of evidence
1 – 3	"barely worth mentioning"
3 - 10	"substantial"
10 - 30	"strong"
> 30	"very strong"

### Model selection

- This quantity is usually difficult to compute, and we can instead use an **approximation** to this ratio
- A common approach is to calculate the **Akaike information criteria** for each model:

$$AIC = \chi_{\min}^2 + 2p + \frac{2p(p+1)}{N-p-1}$$

- p = number of parameters N = number of bins
- This penalizes models with more parameters (and the final term corrects for sample size)
- The model with the smaller value of AIC is preferred [the likelihood ratio is  $e^{(AIC_1 AIC_2)/2}$ ]

## Flat or curved Universe?

- Let's return to the same **supernova distance-redshift dataset** again:
- Compute the Akaike information criteria for a **flat model** (where  $\Omega_m + \Omega_\Lambda =$ 1) and a **curved model** (where  $\Omega_m$ ,  $\Omega_{\Lambda}$  can take any value). Which model is preferred, by this *metric?*



## Summary

At the end of this class you should be able to ...

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